# Notes on the orography variance and slope covariances: filling gaps and separating scales

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#### Abstract

"Separating scales" from databases is recommended when calculating subgrid orographic fields for physical parametrizations. In principle, this could be achieved by pre-pocessing the databases themselves. Unfortunately, these pre-processed databases are not yet available at CMC. These notes are meant to provide an alternative to this pre-processing, by describing a method to approximately (i) fill up the high-wavenumber spectrum of orography and (ii) separate scales of the subgrid variance and slope covariances in a way that is convenient for the boundary layer (PBL) and subgrid orography (SSO) parameterizations.

### 1 Total variance of subgrid orography

Suppose that the orography power spectrum S(K) generally obeys a power law,

$$S(k) = A \cdot K^{-\beta} \tag{1}$$

where A is some amplitude (which varies from a grid-cell to another), K is the total wavenumber, and  $\beta$  is a known exponent. In fact, observations suggest that  $\beta \sim 2$  (see Fig. 1). Then the total variance of the sugbrid orography for a given grid-cell would be

$$\sigma_t^2 = \int_{K_m}^{\infty} S(K) \, dK = \frac{A}{(\beta - 1)} K_m^{-(\beta - 1)} \tag{2}$$

where

$$K_m = \frac{2\pi}{L_m} \tag{3}$$

is the wavenumber associated with the grid-cell size, which is given by the model resolution (i.e.  $L_m \sim \sqrt{\Delta x \Delta y}$ ).

#### 2 Total variance reconstruction from partial variance

An estimate of orography variance is currently produced by the generator of geophysical fields (e.g. Genesis or GenPhysX) for each grid-cell, using elevation data from a database chosen by the user.

Suppose that the chosen database has a resolution  $L_b$  (e.g. GTOPO30 has  $L_b \sim 900m$ ), that  $L_b < L_m$  (i.e. the resolution of database is higher than the model resolution), and let

$$K_b = \frac{2\pi}{L_b} \tag{4}$$

be the associated wavenumber. Then the grid-cell orography variance obtained from this database would be

$$\sigma_b^2 = \int_{K_m}^{K_b} S(k) \, dk = \frac{A}{(\beta - 1)} \left( K_m^{-(\beta - 1)} - K_b^{-(\beta - 1)} \right) \tag{5}$$

Note that  $\sigma_b$  is an underestimation of the total variance, since it lacks contributions from the small scales not resolved by the database.

Still, once  $\sigma_b$  is obtained, we may find (and eliminate) the unknown amplitude A, and so reconstruct the total variance  $\sigma_t$  from  $\sigma_b$  and from the resolution ratio:

$$\sigma_t^2 = \sigma_b^2 \frac{K_m^{-(\beta-1)}}{\left(K_m^{-(\beta-1)} - K_b^{-(\beta-1)}\right)} = \frac{\sigma_b^2}{1 - \left(\frac{L_b}{L_m}\right)^{\beta-1}}$$
(6)

Defining the auxiliary parameter

$$w_{mb} = \left(\frac{L_b}{L_m}\right)^{\beta - 1} \tag{7}$$

we have

$$\sigma_t^2 = \frac{\sigma_b^2}{(1 - w_{mb})} \tag{8}$$

Note 1: The above equations only make sense if the database has enough resolution compared to your model resolution, i.e. if  $L_b < L_m$ . Ideally, you should have  $L_b \ll L_m$ . If not, just get yourself a better database!

#### **3** Separation of total variance

Now suppose that we want to separate the total variance  $\sigma_t$  into a *large*- and a *small*-component, based on a prescribed separation scale  $L_s$  (and the associated wavenumber  $K_s = 2\pi/L_s$ ). This type of separation is in fact useful for the partition of orography forcing between the PBL scheme and the SSO scheme, in which case the recommended separation scale is  $L_s \sim 5km$ .

Assuming that  $L_b < L_s < L_m$ , then we may decompose

$$\sigma_t^2 = \sigma_{large}^2 + \sigma_{small}^2 \tag{9}$$

where

$$\sigma_{small}^{2} = \int_{K_{s}}^{\infty} S(K) \, dK = \frac{A}{(\beta - 1)} K_{s}^{-(\beta - 1)} = \sigma_{t}^{2} \left(\frac{K_{s}}{K_{m}}\right)^{-(\beta - 1)} = \sigma_{t}^{2} \left(\frac{L_{s}}{L_{m}}\right)^{\beta - 1} \tag{10}$$

and

$$\sigma_{large}{}^{2} = \int_{K_{m}}^{K_{s}} S(K) \ dK = \sigma_{t}{}^{2} \left[ 1 - \left(\frac{L_{s}}{L_{m}}\right)^{\beta - 1} \right]$$
(11)

Note that the above separation formula may be generalized to the case when  $L_m < L_s$ (e.g. high-resolution models), if we define the auxiliary separation parameter

$$w_{ms} = \left[\min\left(\frac{L_s}{L_m}, 1\right)\right]^{\beta - 1} \tag{12}$$

so that

$$\sigma_{small}^{2} = w_{ms} \cdot \sigma_{t}^{2} \tag{13}$$

$$\sigma_{large}{}^2 = (1 - w_{ms}) \cdot \sigma_t{}^2 \tag{14}$$

If your model resolution is higher than the separation scale, i.e.  $L_m < L_s$ , then  $w \to 1$  and all the variance goes into the *small*-scale component.

Combining the above relations with those of the previous section, we finally get

$$\sigma_{small}^{2} = \frac{w_{ms}}{(1 - w_{mb})} \cdot \sigma_{b}^{2} \tag{15}$$

$$\sigma_{large}{}^2 = \frac{(1 - w_{ms})}{(1 - w_{mb})} \cdot \sigma_b{}^2 \tag{16}$$

which show how to estimate the requested output (i.e.  $\sigma_{small}^2$  and  $\sigma_{large}^2$ ) from the provided input (i.e.  $\sigma_b^2$  and the resolution ratios).

Note 2: We should probably consider the notion of *effective* resolution, both for the model and for the database, when defining the values of the parameters  $L_m$  and  $L_b$ .

#### 4 Slope covariances

A similar re-scaling may be derived for the slope covariances,

$$G_{xx} = \int_{K_m}^{K_s} k^2 S(K) \, dK \tag{17}$$

$$G_{yy} = \int_{K_m}^{K_s} l^2 S(K) \, dK \tag{18}$$

$$G_{xy} = \int_{K_m}^{K_s} k l S(K) \ dK \tag{19}$$

where k and l indicate wavenumbers in the x- and y- directions respectively, and  $K^2 = k^2 + l^2$ . Note that the SSO scheme only needs the *large*-scale component of these quantities, therefore the integral limits  $K_m$  to  $K_s$ .

Assuming approximate isotropy, i.e. that  $G_{xx} \sim G_{yy} \sim 0.5 \int (k^2 + l^2) S(K) \, dK$ , we could estimate

$$G_{xx} \sim 0.5 \int_{K_m}^{K_s} K^2 S(K) \ dK = 0.5 \frac{A}{(3-\beta)} \left( K_s^{(3-\beta)} - K_m^{(3-\beta)} \right)$$
(20)

If we use all the wavelengths available in the high-resolution database to compute the covariance, we will be oversestimating this quantity by

$$G_{xx}^{b} \sim 0.5 \int_{K_{m}}^{K_{b}} K^{2}S(K) \ dK = 0.5 \frac{A}{(3-\beta)} \left( K_{b}^{(3-\beta)} - K_{m}^{(3-\beta)} \right) \sim \frac{1}{r} G_{xx}$$
(21)

where the factor r is given by

$$r = \frac{\left(K_s^{(3-\beta)} - K_m^{(3-\beta)}\right)}{\left(K_b^{(3-\beta)} - K_m^{(3-\beta)}\right)} = \frac{\left(\frac{L_m}{L_s}\right)^{(3-\beta)} - 1}{\left(\frac{L_m}{L_b}\right)^{(3-\beta)} - 1}$$
(22)

Note that, if  $\beta \sim 2$ , then 0 < r < 1 whenever  $L_b < L_s < L_m$ , consistent with the statement that  $G_{xx}^b$  is an overestimate of  $G_{xx}$ .

Assuming that  $G_{xx}^b$  is the quantity provided by GenPhysX, then the adjusted (re-scaled) value we actually want would be

$$G_{xx} \sim r \cdot G_{xx}^b \tag{23}$$

We could use the same re-scaling to adjust  $G_{yy}$  and  $G_{xy}$ .

Note 3: Alternatively (ideally), we would avoid the need of such re-scaling if we could compute the covariances directly from a database of resolution  $L_b \sim L_s$ .

#### 5 Generalization to the case of a 2-exponent spectrum

Some studies suggest that the orography spectrum may actually have the form:

$$S(K) = A \cdot \begin{cases} K^{-\beta_1}, & \text{if } K \le K_0 \\ K_0^{(\beta_2 - \beta_1)} K^{-\beta_2}, & \text{if } K > K_0 \end{cases}$$
(24)

where  $\beta_1$  and  $\beta_2$  are distinct exponents, and  $K_0$  is a the wavenumber where the that change occurs. For instance, Beljaars et al. 2004 propose  $\beta_1 = 1.9$ ,  $\beta_2 = 2.8$ , and  $K_0 = 0.003m^{-1}$  (i.e.  $L_0 \sim 2km$ ).

Hereafter we will assume that the database has a sufficiently high resolution, i.e. that  $L_b < L_0$  and  $L_b < L_m$ . In this case, the calculation of the total variance splits into 2 cases: (i) if  $L_m > L_0$  ( $K_m < K_0$ ):

$$\sigma_t^2 = \int_{K_m}^{\infty} S(K) \, dK \tag{25}$$

$$= A \cdot \left[ \frac{1}{(\beta_1 - 1)} K_m^{-(\beta_1 - 1)} + \frac{(\beta_1 - \beta_2)}{(\beta_1 - 1)(\beta_2 - 1)} K_0^{-(\beta_1 - 1)} \right]$$
(26)

(ii) if  $L_m < L_0 \ (K_m > K_0)$ :

$$\sigma_t^2 = \int_{K_m}^{\infty} S(K) \ dK = A \cdot \left[ \frac{K_0^{(\beta_2 - \beta_1)}}{(\beta_2 - 1)} K_m^{-(\beta_2 - 1)} \right]$$
(27)

Meanwhile, the calculation of  $\sigma_b^2$  also falls into 2 cases: (i) if  $L_m > L_0$  ( $K_m < K_0$ ):

$$\sigma_b^2 = \int_{K_m}^{K_b} S(K) \, dK \tag{28}$$

$$= A \cdot \left[ \frac{1}{(\beta_1 - 1)} K_m^{-(\beta_1 - 1)} + \frac{(\beta_1 - \beta_2)}{(\beta_1 - 1)(\beta_2 - 1)} K_0^{-(\beta_1 - 1)} - \frac{K_0^{(\beta_2 - \beta_1)}}{(\beta_2 - 1)} K_b^{-(\beta_2 - 1)} \right] (29)$$

$$= \sigma_t^2 \cdot \left[ 1 - \frac{1}{c} \left( \frac{L_b}{L_m} \right)^{(\beta_2 - 1)} \right]$$
(30)

where

$$c = \frac{(\beta_2 - 1)}{(\beta_1 - 1)} \left(\frac{L_0}{L_m}\right)^{(\beta_2 - \beta_1)} + \frac{(\beta_1 - \beta_2)}{(\beta_1 - 1)} \left(\frac{L_0}{L_m}\right)^{(\beta_2 - 1)}$$
(31)

(ii) if  $L_m < L_0 \ (K_m > K_0)$ :

$$\sigma_b^2 = \int_{K_m}^{K_b} S(K) \, dK = A \cdot \left[ \frac{K_0^{(\beta_2 - \beta_1)}}{(\beta_2 - 1)} \left( K_m^{-(\beta_2 - 1)} - K_b^{-(\beta_2 - 1)} \right) \right]$$
(32)

$$= \sigma_t^2 \cdot \left[ 1 - \left(\frac{L_b}{L_m}\right)^{(\beta_2 - 1)} \right] \tag{33}$$

Note that the 2 cases may now be unified as follows:

$$\sigma_t^2 = \frac{\sigma_b^2}{(1 - w_{mb})} \tag{34}$$

where

$$w_{mb} = \frac{1}{c} \left(\frac{L_b}{L_m}\right)^{(\beta_2 - 1)} \tag{35}$$

$$c = \frac{(\beta_2 - 1)}{(\beta_1 - 1)} a^{(\beta_2 - \beta_1)} + \frac{(\beta_1 - \beta_2)}{(\beta_1 - 1)} a^{(\beta_2 - 1)}$$
(36)

$$a = \min\left(\frac{L_0}{L_m}, 1\right) \tag{37}$$

As for the separation of scales, we must consider 2 different cases:

(i) if  $L_m > L_s$  ( $K_m < K_s$ ), i.e. for a relatively course-resolution model:

$$\sigma_{small}^2 = \int_{K_s}^{\infty} S(K) \, dK \tag{38}$$

$$= A \cdot \left[ \frac{1}{(\beta_1 - 1)} K_s^{-(\beta_1 - 1)} + \frac{(\beta_1 - \beta_2)}{(\beta_1 - 1)(\beta_2 - 1)} K_0^{-(\beta_1 - 1)} \right]$$
(39)

$$= \sigma_t^2 \cdot \left(\frac{L_s}{L_m}\right)^{(\beta_1 - 1)} \frac{\left[1 + \frac{(\beta_1 - \beta_2)}{(\beta_2 - 1)} \left(\frac{L_0}{L_s}\right)^{(\beta_1 - 1)}\right]}{\left[1 + \frac{(\beta_1 - \beta_2)}{(\beta_2 - 1)} \left(\frac{L_0}{L_s}\right)^{(\beta_1 - 1)} \left(\frac{L_s}{L_m}\right)^{(\beta_1 - 1)}\right]}$$
(40)

(ii) if  $L_m < L_s$  ( $K_m > K_s$ ), i.e. for a relatively high-resolution model:

$$\sigma_{small}^2 = \sigma_t^2 \tag{41}$$

The relations above may also be unified as follows:

$$\sigma_{small}^2 = w_{ms} \cdot \sigma_t^2 \tag{42}$$

where

$$w_{ms} = b^{(\beta_1 - 1)} \frac{\left[1 + \frac{(\beta_1 - \beta_2)}{(\beta_2 - 1)} \left(\frac{L_0}{L_s}\right)^{(\beta_1 - 1)}\right]}{\left[1 + \frac{(\beta_1 - \beta_2)}{(\beta_2 - 1)} \left(\frac{L_0}{L_s}\right)^{(\beta_1 - 1)} b^{(\beta_1 - 1)}\right]}$$
(43)

and

$$b = \min\left(\frac{L_s}{L_m}, 1\right) \tag{44}$$

In sum, the relations

$$\sigma_{small}^2 = \frac{w_{ms}}{(1 - w_{mb})} \cdot \sigma_b^2 \tag{45}$$

$$\sigma_{large}^2 = \frac{(1 - w_{ms})}{(1 - w_{mb})} \cdot \sigma_b^2 \tag{46}$$

(47)

are still valid, all we need to use is the generalized formulas for  $w_{ms}$  and  $w_{mb}$ .

The same type of generalization may be extended to the scaling of slope covariances. Except that now we only care for the case in which  $L_m > L_s$  ( $K_m < K_s$ ):

$$G_{xx} \sim 0.5 \int_{K_m}^{K_s} K^2 S(K) \, dK = 0.5 \frac{A}{(3-\beta_1)} \left( K_s^{(3-\beta_1)} - K_m^{(3-\beta_1)} \right) \tag{48}$$

and

$$G_{xx}^b \sim 0.5 \int_{K_m}^{K_b} K^2 S(K) \, dK$$
 (49)

$$= 0.5 \frac{A}{(3-\beta_1)} \left[ \frac{(3-\beta_1)}{(3-\beta_2)} \left( \frac{K_0}{K_b} \right)^{(\beta_2-\beta_1)} K_b^{(3-\beta_1)} - K_m^{(3-\beta_1)} \right]$$
(50)

+ 
$$\frac{(\beta_1 - \beta_2)}{(3 - \beta - 1)(3 - \beta_2)} K_0^{(3 - \beta_1)}$$
 (51)

which implies that

$$G_{xx} \sim r \cdot G^b_{xx} \tag{52}$$

where the scaling factor now reads

$$r = \frac{1 - b^{(3-\beta_1)}}{\frac{(3-\beta_1)}{(3-\beta_2)} \left(\frac{L_b}{L_0}\right)^{(\beta_2-\beta_1)} \left(\frac{L_s}{L_b}\right)^{(3-\beta_1)} - b^{(3-\beta_1)} + \frac{(\beta_1-\beta_2)}{(3-\beta_1)(3-\beta_2)} \left(\frac{L_s}{L_0}\right)^{(3-\beta_1)}}$$
(53)

Examples of scaling factors for a range of model resolutions are shown in Figure 2.

## 6 Implementation in GenPhysX

Only the formulas based on the 1-exponent spectrum are currently implemented in Gen-PhysX. To activate this scale-separation method in GenPhysX, it suffices to set the option

-subgrid SPLIT

The default values of the spectrum exponent  $\beta$  and the separation wavelength  $L_s$  (in m) are set, respectively, as

set Const(beta) 2. set Const(lres) 5000.0

The values of these parameteres may in principle be re-set by the user.



Figure 1: Power spectrum of orography derived from the resolved (non-filtered) topography elevation field of the GDPS-25km model. In the x-axis, N is what was indicated as K in the formulas above. A spectrum with  $\beta = 2$  is also shown for reference.



Figure 2: Scaling factors for variances and slope covariances for various model resolutions, assuming a database with resolution  $L_b = 0.9km$ . Solid lines correspond to the 2-exponent slope assumption  $\beta_1 = 1.9, \beta_2 = 2.8$ , and dashed lines correspond to the single slope assumption ( $\beta_1 = \beta_2 = 2$ ).